

SOLUTIONS & ADDITIONAL REMARKS FOR LAB #2

1. $L(y_1, \dots, y_n; \theta) = (2T)^{-n} e^{-\frac{1}{T} \sum |y_i - \theta|}$; to maximize L we must minimize $\sum |y_i - \theta|$. We can use calculus, as follows:
- $$\frac{d}{d\theta} \sum |y_i - \theta| = \begin{cases} y_i - \theta & \text{if } y_i > \theta \\ \theta - y_i & \text{if } y_i < \theta \end{cases}; \text{ hence } \frac{d}{d\theta} |y_i - \theta| = \begin{cases} -1 & \text{if } y_i > \theta \\ 1 & \text{if } y_i < \theta \end{cases}.$$

Thus $\frac{d}{d\theta} \sum |y_i - \theta| = \sum_{i: y_i > \theta} 1 + \sum_{i: y_i < \theta} (-1) = \#(y_i > \theta) - \#(y_i < \theta)$.

This derivative will be 0 precisely when θ is the (a) median, by the definition of the median.

5. $\ln f(y; \theta) = -\ln(2T) - \frac{1}{T} |y - \theta|$; hence $\frac{d \ln f}{d\theta} = \begin{cases} -1/T & \text{if } y > \theta \\ 1/T & \text{if } y < \theta \end{cases}$
Hence $I(\theta) = \frac{1}{n E\left(\frac{d \ln f}{d\theta}\right)^2} = \frac{1}{n \left(\frac{1}{T^2}\right)} = T^2/n$, which = $\text{Var}(\tilde{Y})$ (see #2 below).

(Note: $\frac{d \ln f}{d\theta}$ does not exist at $y = \theta$; however, since $P(Y = \theta) = 0$ since Y is a continuous r.v., this has no effect on $E\left(\frac{d \ln f}{d\theta}\right)^2$.)

2. $\text{Var } \bar{Y} = \frac{\sigma^2}{n} = \frac{1}{n} E(Y - \mu)^2 = \frac{1}{n} \int_{-\infty}^{\infty} (y - \theta)^2 \cdot \frac{1}{2T} e^{-\frac{1}{T} |y - \theta|} dy$ (since $\mu = \theta$)
 $= \frac{2}{n} \int_0^{\infty} \dots$ since the integrand is a symmetric (even) function;
let $u = \frac{y - \theta}{T}$, so $du = \frac{dy}{T}$, $\Rightarrow \text{Var } \bar{Y} = \frac{2}{n} \int_0^{\infty} (Tu)^2 \cdot \frac{1}{2T} e^{-|u|} \cdot T du$
 $= \frac{T^2}{n} \int_0^{\infty} u^2 e^{-u} du = \frac{T^2}{n} \Gamma(3) = \frac{T^2}{n} \cdot 2! = \frac{2T^2}{n}$ (recall the gamma fn., p. 165)

$$\text{Var } \tilde{Y} = \frac{1}{4n f^2(\theta)}, \text{ where } f(\theta) \text{ means } f(y; \theta) \text{ evaluated at } y = \theta,$$

$$= \frac{1}{4n \cdot \left(\frac{1}{2T} e^{-\frac{1}{T} |\theta - \theta|}\right)^2} = \frac{1}{4n \cdot (1/2T)^2} = \frac{T^2}{n}.$$

3. $\text{Eff}(\tilde{Y}, \bar{Y}) = \text{Var } \bar{Y} / \text{Var } \tilde{Y} = \frac{2T^2/n}{T^2/n} = 2$. The mean requires twice as much data to do as well as the median here. A consequence is that \tilde{Y} yields narrower CI's (see #4).

4. We use the formula $\hat{\theta} \pm 2 \hat{\sigma}_\theta$ to obtain

$$\bar{Y} \pm 2 \sqrt{\frac{2\hat{\tau}^2}{n}} = 110.5 \pm 2 \sqrt{\frac{2(100)^2}{15}} = 110.5 \pm 73.0$$

$$\text{and } \tilde{Y} \pm 2 \sqrt{\frac{\hat{\tau}^2}{n}} = 91 \pm 2 \sqrt{\frac{100^2}{15}} = 91 \pm 51.6.$$

As expected, the interval based on \tilde{Y} is narrower.

5. $\ln f(y; \theta) = -\ln(2\tau) - \frac{1}{\tau}|y - \theta|$, thus

$$\frac{\partial \ln f(y; \theta)}{\partial \theta} = \begin{cases} 1/\tau & \text{if } \theta < y, \\ -1/\tau & \text{if } \theta > y. \end{cases}$$

$$\text{Thus } I(\theta) = \frac{1}{n E\left[\left(\frac{\partial \ln f}{\partial \theta}\right)^2\right]} = \frac{1}{n E\left[\left(\pm \frac{1}{\tau}\right)^2\right]} = \frac{1}{n \left(\frac{1}{\tau}\right)^2} = \frac{\tau^2}{n}.$$

Thus \tilde{Y} does achieve the Cramer-Rao lower bound.
Hence it is the MVUE of θ .

$$6. L(\tau) = \prod_{i=1}^n (2\tau)^{-1} e^{-\frac{1}{\tau}|y_i - \theta|} = (2\tau)^{-n} e^{-\frac{1}{\tau} \sum |y_i - \theta|}$$

$$\Rightarrow \ln L(\tau) = -n \ln(2\tau) - \frac{1}{\tau} \sum |y_i - \theta|$$

$$\Rightarrow \frac{\partial \ln L(\tau)}{\partial \tau} = -\frac{n}{\tau} + \frac{1}{\tau^2} \sum |y_i - \theta| \stackrel{!}{=} 0 \Rightarrow \hat{\tau} = \frac{\sum |y_i - \theta|}{n}.$$

Since θ is unknown we must use its m.l.e. in place of θ to obtain

$$\begin{aligned} \hat{\tau} &= \frac{\sum |y_i - \tilde{y}|}{n} = \frac{1}{15} \left[(209 - 91) + (341 - 91) + \dots + (-13 - 91) \right] \\ &= \frac{1}{15} (1500) = 100. \end{aligned}$$